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## ON MOMENTS OF PROBABILITY DISTRIBUTION FUNCTIONS AND H-FUNCTION TRANSFORM

### ABSTRACT

We express the absolute moments of a probability distribution function in terms of the H-function transform. A special case and an example is cited to illustrate the application of our general result.

### RESUMEN

En este trabajo se obtienen momentos absolutos de una función de distribución en términos de la función H. Se considera un caso especial y un ejemplo para ilustrar la aplicación de la fórmula general.

### 1. INTRODUCTION

If  $f(x)$  is a probability distribution function and  $\phi(p)$  is the Laplace-Stieltjes transform defined by

$$(1.1) \quad \phi(p) = \int_0^\infty e^{-xp} df(x), \quad p \geq 0,$$

then  $\phi(p)$  possesses derivatives of all orders given by

$$(1.2) \quad \int_0^\infty e^{-xp} x^n df(x) = (-1)^n \phi^{(n)}(p), \quad 0 \leq p < \infty,$$

where  $f(x) = 0$  for  $x < 0$ .

It follows in particular (see Feller [2]) that the probability distribution function  $f(x)$  has a finite absolute moment of the  $n^{\text{th}}$  order iff  $\phi^{(n)}(0)$  exists such that

$$(1.3) \quad \int_0^\infty |x|^n df(x) = (-1)^n \phi^{(n)}(0).$$

Denote by  $A$  the class of all such functions  $f$  which are differentiable everywhere any number of times and if it and all of its derivatives are  $O(x^{-v})$ , for all  $v$  as  $x$  increases without limit.

The Weyl fractional derivatives of a function  $g(x)$  is defined as follows :

$$(1.4) \quad {}_{x^0} D_x^\gamma g(x) = \frac{(-1)^\gamma}{\Gamma(-\gamma)} \int_x^\infty (t-x)^{-\gamma-1} g(t) dt, \quad \text{for } \gamma < 0$$

For  $\gamma \geq 0$ ,

$$(1.5) \quad {}_{x^0} D_x^\gamma g(x) = \frac{d^r}{dx^r} ({}_{x^0} D_x^{\gamma-r} g(x)),$$

where  $r$  is a positive integer such that  $r-\gamma > 0$ . Whenever  $g \in A$ , the representations (1.4) and (1.5) exist (see Miller [5, p. 80]).

In a recent paper Wolfe [11] extended the law of moments of order  $n$  ( $n$  a positive integer) to an equivalent result valid for an arbitrary order  $\lambda$  ( $\lambda$  real). We find it convenient to record this result in the following form :

**THEOREM 1** (Wolfe [11]). Let  $f(x)$  be a probability distribution function such that  $f(x)=0$  for  $x<0$ . Then  $f(x)$  possesses an absolute moment of the  $\lambda^{\text{th}}$  order iff  $p D_x^\lambda g(0)$  exists such that

$$(1.6) \quad \int_0^\infty x^\lambda df(x) = (-1)^{-\lambda} p D_x^\lambda g(0), \quad \lambda \text{ real},$$

where  $g(p)$  denotes the Laplace-Stieltjes transform as defined in (1.1), of  $f(x)$ .

**REMARK 1.** It may be observed that the proof given by Wolfe [11, p.310] can also be formulated with the aid of an interesting relation giving the Laplace transform of the function  $t^\gamma f(t)$ , for all real  $\gamma$ , via Weyl fractional calculus, given in the paper of Raina and Koul [7, p.189].

Our object in this paper is to express the absolute moments of a probability distribution function in terms of the H-function transform, defined below, which provides a generalization of the aforesaid Theorem 1. An example is cited to illustrate the application of our generalized result.

## 2. THE MAIN RESULT

If  $f(x)$  is the probability distribution function such that  $f(x) = 0$ , for  $x < 0$ , then the integral equation

$$(2.1) \quad g(p) = \int_0^\infty H_{P,Q}^{M,N} \left[ a(pt)^h \middle| \begin{matrix} (a_i, \alpha_i)_{1,P} \\ (b_i, \beta_i)_{1,Q} \end{matrix} \right]$$

$$d f(t) dt, \quad h > 0, \quad 0 \leq p < \infty,$$

defines an H-function transform of  $f(t)$ .

In particular, if

$$(2.2) \quad f(t) = \int_0^t \phi(x) dx, \quad 0 \leq t < \infty,$$

then (2.1) reduces to the equivalent form of the H-function transform due to Gupta and Mittal [4, p. 142].

The H-function Kernel in (2.1) is the well known H-function of C. Fox [1, p.408] which we define in the following form :

$$(2.3) \quad H_{P,Q}^{M,N} \left[ z \middle| \begin{matrix} (a_i, \alpha_i)_{1,P} \\ (b_i, \beta_i)_{1,Q} \end{matrix} \right] = \frac{1}{2\pi w} \int_L \theta(s) z^s ds,$$

with  $w = \sqrt{-1}$ , and  $\theta(s)$  given by

$$(2.4) \quad \theta(s) = \left( \prod_{j=1}^M \Gamma(b_j - \beta_j s) \right) \left( \prod_{j=1}^N \Gamma(1 - a_j + \alpha_j s) \right) \left( \prod_{j=M+1}^Q \Gamma(1 - b_j + \beta_j s) \right) \left( \prod_{j=N+1}^P \Gamma(a_j - \alpha_j s) \right)^{-1}.$$

The contour L is a suitably chosen contour of Mellin-Barnes type; the non negative integers M, N, P, Q satisfy the inequalities  $0 \leq M \leq Q$ ,  $0 \leq N \leq P$ . The parameters  $\alpha_i, \beta_i$  are all positive and the parameters  $a_i, b_i$  are arbitrary complex numbers.

It being understood that  $(a_i, \alpha_i)_{1,P}$  condenses the array of parameters  $(a_1, \alpha_1), \dots, (a_p, \alpha_p)$ ,  $p \geq 0$ ; with similar interpretation for  $(b_i, \beta_i)_{1,Q}$ . For a detailed account of the conditions of existence of the H-function (2.3), its various special cases and other important properties, we refer to the paper of Gupta and Jain [3].

Before stating our main result of this paper which provides a generalization of (1.6), we require the following interesting result :

LEMMA. Let  $f(x)$  be a probability distribution function whose H-function transform  $g(p)$  given by (2.1.) be such that  $g(p) \in A$ . Suppose

(i)  $0 \leq M \leq Q$ ,  $0 \leq N \leq P$  ( $M, N, P, Q$  being non-negative integers),

(ii)  $h[(a_i - 1)/\alpha_i] < 0$  ( $i = 1, \dots, N$ ),  $h > 0$ ,  $0 \leq p < \infty$ ,

(iii)  $|\arg(a)| < \frac{1}{2} \Delta \pi$ ,  $\Delta = \sum_{i=1}^M \beta_i - \sum_{i=1}^N \alpha_i - \sum_{i=1}^P \alpha_i > 0$ .

Then  $g(p)$  possesses derivatives of arbitrary order  $\gamma$  (real) given by

$$(2.5) \quad a(-1)^{-\gamma} p \int_0^\infty H_{P+1,Q+1}^{M+1,N} \left[ x^\gamma f(x); p \right] dt,$$

where

$$(2.6) \quad \psi[x^\gamma f(x); p] = \int_0^\infty H_{P+1,Q+1}^{M+1,N} \left[ a(px)^h \left| \begin{matrix} (a_i - \gamma h, \alpha_i)_{1,P}, (-\gamma, h) \\ (0, h), (b_i - \frac{\gamma \beta_i}{h}, \beta_i)_{1,Q} \end{matrix} \right. \right] x^\gamma df(x).$$

Proof : If  $f(x)$  is a probability distribution function such that  $f(x) = 0$  for  $x < 0$ , then for  $\gamma < 0$ , we have in view of (1.4) and (2.1),

$$\begin{aligned} p \int_0^\infty H_{P+1,Q+1}^{M+1,N} \left[ a(-1)^{-\gamma} \int_p^\infty (t-p)^{-\gamma-1} g(t) dt \right] dt \\ = \frac{(-1)^\gamma}{\Gamma(-\gamma)} \int_p^\infty (t-p)^{-\gamma-1} \left( \int_0^\infty H_{P,Q}^{M,N} \left[ a(tx)^h \left| \begin{matrix} (a_i, \alpha_i)_{1,P} \\ (b_i, \beta_i)_{1,Q} \end{matrix} \right. \right] dt \right) df(x) dt \\ = \frac{(-1)^\gamma}{\Gamma(-\gamma)} \int_0^\infty \left( \int_0^\infty (t-p)^{-\gamma-1} H_{P,Q}^{M,N} \left[ a(tx)^h \left| \begin{matrix} (a_i, \alpha_i)_{1,P} \\ (b_i, \beta_i)_{1,Q} \end{matrix} \right. \right] dt \right) df(x). \end{aligned}$$

Since

$$(2.7) \quad \int_p^\infty (x-p)^{-\gamma-1} x^\mu H_{P+1,Q+1}^{M+1,N} \left[ a x^h \left| \begin{matrix} (a_i, \alpha_i)_{1,P}, (-\mu, h) \\ (\gamma-\mu, h), (b_i, \beta_i)_{1,Q} \end{matrix} \right. \right] dx \\ = p^{\mu-\gamma} \Gamma(-\gamma) H_{P+1,Q+1}^{M+1,N} \left[ a p^h \left| \begin{matrix} (a_i, \alpha_i)_{1,P}, (-\mu, h) \\ (\gamma-\mu, h), (b_i, \beta_i)_{1,Q} \end{matrix} \right. \right]$$

provided that  $h > 0$ ,  $\operatorname{Re}[h(a_i - 1)/\alpha_i] < \operatorname{Re}(\gamma) < 0$  ( $i = 1, \dots, N$ ),  $|\arg(a)| < \frac{1}{2} \Delta \pi$ ,  $\Delta$  being given in condition (iii) (stated with Theorem 2 above). (which

can rather easily be established by invoking the definitions (1.4) and (2.3); also deducible from the result given by Raina [9, p.40, Eq.(3.3)], we find that

$$(2.8) \quad p_{\infty}^{D^{\gamma}} g(p) = (-1)^{\gamma} \int_0^{\infty} x^{\gamma} H_{p+1, Q+1}^{M+1, N}$$

$$\left[ a(px)^h \begin{bmatrix} (a_i, \alpha_i)_{1,p}, (0, h) \\ (\gamma, h), (b_i, \beta_i)_{1,Q} \end{bmatrix} \right] df(x)$$

$$= \frac{(-1)^{\gamma}}{a} \int_0^{\infty} x^{\gamma} H_{p+1, Q+1}^{M+1, N}$$

$$\left[ a(px)^h \begin{bmatrix} (a_i - q\alpha_i/h, \alpha_i)_{1,p}, (-\gamma, h) \\ (0, h), (b_i - q\beta_i/h, \beta_i)_{1,Q} \end{bmatrix} \right] x df(x),$$

which evidently establishes our result for  $\gamma < 0$ .

For  $\gamma \geq 0$ , by invoking the definition (1.5) and the case (2.8), we get

$$(2.9) \quad p_{\infty}^{D^{\gamma}} g(p) = \frac{(-1)^{\gamma-n}}{a} \frac{d^n}{dp^n} \left( \int_0^{\infty} x^{\gamma} H_{p+1, Q+1}^{M+1, N} \right)$$

$$\left[ a(px)^h \begin{bmatrix} (a_i - q\frac{\alpha_i}{h}, \alpha_i)_{1,p}, (-q, h) \\ (0, h), (b_i - q\frac{\beta_i}{h}, \beta_i)_{1,Q} \end{bmatrix} \right] df(x),$$

where  $\gamma' = \gamma - n$

Differentiating under the sign of integral (justifiable under the conditions stated with the theorem and the Lebesgue dominated convergence theorem), appealing to the formula of Skibiniski [10, p. 131, Eq.(4.1)] regarding the derivatives of  $H$ -function (see also Raina and Koul [6]), we arrive at the required result (2.5) for  $\gamma \geq 0$ . This proves our lemma.

**REMARK 2.** It must be mentioned that the aforesaid lemma would also follow from a very recently established result due to Raina and Koul [8], with of course, suitable amendments in the hypothesis and derivation method.

**THEOREM 2.** If  $f(x)$  is a probability distribution function such that  $f(x) = 0$  for  $x < 0$  whose  $H$ -function transform  $g(p)$  given by (2.1) exist. Then  $g(p)$  possesses an absolute moment of the  $\gamma$ th order

( $-\infty < \gamma < \infty$ ), iff  $p_{\infty}^{D^{\gamma}}$  exists and is given by

$$(2.10) \quad \int_0^{\infty} x^{\gamma} df(x) = a(-1)^{-\gamma} \prod_{i=1}^Q \frac{\Gamma(1-b_i+\gamma)}{p} p_{\infty}^{D^{\gamma}} g(0),$$

where

$$(2.11) \quad \prod_{i=1}^Q \frac{\Gamma(1-b_i+\gamma)}{p} = \frac{\prod_{i=1}^Q \Gamma(1-a_i+\gamma)}{\Gamma(1+\gamma) \prod_{i=1}^Q \Gamma(1-a_i+\gamma)},$$

$$|\arg(a)| < \frac{\pi(P-Q+1)}{2},$$

$a_i, b_i$  are arbitrary complex parameters and assume values such that the Gamma quotients in (2.11) exist

**Proof :** suppose  $f(x)$  is a probability distribution function and let  $f(x)$  possess the absolute moment of the  $\gamma$ th order ( $-\infty < \gamma < \infty$ ).

From the elementary relationship of the  $H$ -function [3, p. 600 Eq. (4.6)]

$$(2.12) \quad \frac{H_{p, Q+1}^{1, P}}{z} \left[ z \begin{bmatrix} (a_i, 1)_{1,p} \\ (b_i, 1)_{1,Q} \end{bmatrix} \right] = \frac{\prod_{i=1}^P \Gamma(a_i)}{\prod_{i=1}^Q \Gamma(b_i)} p^F_Q [(a_p); (b_Q); -z],$$

it follows in particular from our lemma (by setting the parameters in accordance with the relation (2.12) and then finally letting  $p \rightarrow 0$  in the resulting expression) that

$$(2.13) \quad p_{\infty}^{D^{\gamma}} g(0) = \frac{(-1)^{\gamma}}{a} \cdot$$

$$\frac{\prod_{i=1}^P \Gamma(1-a_i+\gamma)}{\prod_{i=1}^Q \Gamma(1-b_i+\gamma)} \int_0^{\infty} x^{\gamma} df(x),$$

which is our desired result (2.10).

Conversely, if  $p_{\infty}^{D^{\gamma}} g(0)$  given by (2.10) exist, then it follows straight forwardly from the Fubini's theorem that

$$\int_0^{\infty} |x|^{\gamma} d f(x) < \infty, \quad \gamma \in (-\infty, \infty),$$

and thus  $f(x)$  possesses the absolute moment of  $\gamma$ th order.

REMARK 3. By the elementary relation

$$(2.14) \quad H_{0,1}^{1,0} \left[ z \mid \frac{1}{(0,1)} \right] = e^{-z},$$

the H-function transform (2.1) reduces to the Laplace-Stieltjes transform (1.1), and the relation (2.11) transforms into

$$(2.15) \quad H_0(-,0;\gamma) = 1.$$

In view of (2.14) and (2.15), Wolfe's result (1.6) follows at once from our Theorem 2.

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